

Quantile Processes for Semi and Nonparametric Regression

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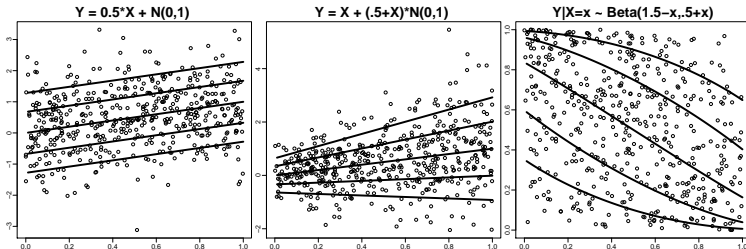
A joint work with Stanislav Volgushev and Guang Cheng

Quantile

Response Y , predictors X . Conditional quantile curve $Q(\cdot; \tau)$ of $Y \in \mathbb{R}$ conditional on X is defined through

$$P(Y \leq Q(X; \tau) | X = x) = \tau \quad \forall x.$$

Examples:



Quantile regression vs mean regression

Mean regression:

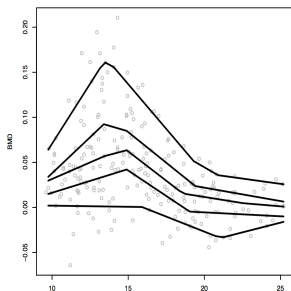
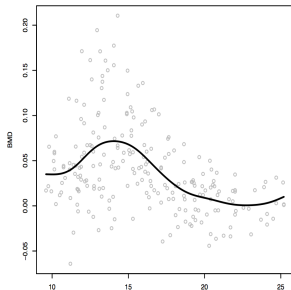
$$Y_i = m(X_i) + \varepsilon_i, \mathbb{E}[\varepsilon|X = x] = 0$$

- m : Regression function, object of interest.
- ε_i : 'errors'.

Quantile regression:

$$P(Y \leq Q(x; \tau) | X = x) = \tau$$

- No strict distinction between 'signal' and 'noise'.
- Object of interest: properties of conditional distribution of $Y|X = x$.
- Contains much richer information than just conditional mean.



Quantile Regression: Estimation

Koenker and Bassett (1978): if $Q(x; \tau) = \beta(\tau)^\top x$, estimate by

$$\hat{\beta}(\tau) := \arg \min_{\mathbf{b}} \sum_i \rho_\tau(Y_i - \mathbf{b}^\top X_i) \quad (1.1)$$

where $\rho_\tau(u) := \tau u^+ + (1 - \tau)u^-$ 'check function'. Well-behaved convex optimization.

$$\hat{Q}(x_0; \tau) := x_0^\top \hat{\beta}(\tau) \text{ for any } x_0.$$

▸ Series Model

▸ Partial Linear Model

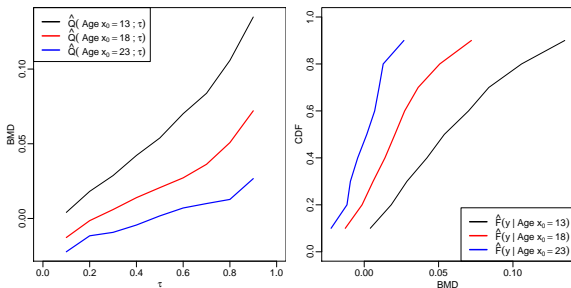
Quantile Regression Process

The **quantile regression process (QRP)** at x_0 is

$$a_n(\widehat{Q}(x_0; \tau) - Q(x_0; \tau)) \in \ell^\infty(\mathcal{T}), \quad (1.2)$$

where $a_n \rightarrow \infty$ appropriately chosen, $\mathcal{T} \subset (0, 1)$ is compact
 For fixed x_0 , and τ_L, τ_U "close" to 0 and 1,

$$Q^{-1}(x_0; y) = F(y|x_0) \approx \tau_L + \int_{\tau_L}^{\tau_U} \mathbf{1}\{Q(x_0; \tau) < y\} d\tau =: \Phi_y(Q(x_0; \tau))$$



Convergence of QRP

If for $a_n \rightarrow \infty$,

$$a_n(\widehat{Q}(x_0; \cdot) - Q(x_0; \cdot)) \rightsquigarrow \mathbb{G}(\cdot) \text{ in } \ell^\infty(\mathcal{T}), \quad (1.3)$$

where $\mathbb{G}(\cdot)$: Gaussian process, $\ell^\infty(\mathcal{T})$: set of all uniformly bounded, real functions on \mathcal{T} , then,

$$a_n\{\widehat{F}_{Y|X}(\cdot|x_0) - F_{Y|X}(\cdot|x_0)\} \rightsquigarrow \\ - f_{Y|X}(\cdot|x_0)\mathbb{G}(x_0; F_{Y|X}(\cdot|x_0)) \text{ in } \ell^\infty(\mathcal{Y}). \quad (1.4)$$

Proof: $\Phi_y(Q(x; \tau))$ is Hadamard differentiable (Chernozhukov et al., 2010) (tangentially to $\mathcal{C}(0, 1)$ at any strictly increasing, differentiable function); functional delta method

Study (1.3):

- Fixed dimension linear model $Q(x; \tau) = x^\top \beta(\tau)$: Koenker and Xiao (2002); Angrist et al. (2006)
- Kernel nonparametric estimation: Qu and Yoon (2015)

A Unified framework (Belloni et al., 2011):

$$Q(x; \tau) \approx \mathbf{Z}(x)^\top \beta(\tau)$$

$\mathbf{Z}(X_i)$ can be a higher dimensional ($\rightarrow \infty$) transformation

- $\hat{Q}(x; \tau) = \mathbf{Z}(x)^\top \hat{\beta}(\tau)$, where $\hat{\beta}(\tau)$ is estimated by replacing X_i by $\mathbf{Z}(X_i)$ in [▶ Quantile Regression](#)
- Need to control the bias $Q(x; \tau) - \mathbf{Z}(x)^\top \beta(\tau)$

Overview

We present process convergence results

$$a_n(\widehat{Q}(x_0; \cdot) - Q(x_0; \cdot)) \rightsquigarrow \mathbb{G}(\cdot) \text{ in } \ell^\infty(\mathcal{T})$$

for the models:

- General series estimator
- B-splines: $\mathbf{Z}(x) = \mathbf{B}(x)$
- An application: partial linear models
 $\mathbf{Z} = (V^\top, \widetilde{\mathbf{Z}}(W)^\top)^\top \in \mathbb{R}^{k+k'}$

Notation:

- $\mathbf{Z}_i := \mathbf{Z}(X_i)$ general basis function (e.g. trigonometric, power, etc.);
- $\mathbf{B}_i := \mathbf{B}(X_i)$ local basis (e.g. B-spline)

Technical Assumptions

Assumption (A): data $(X_i, Y_i)_{i=1, \dots, N}$ form triangular array and are row-wise i.i.d. with

(A1) $m = \mathbf{Z}(x)$. Assume that $\|\mathbf{Z}_i\| \leq \xi_m < \infty$, and there exists some fixed constant M so that

$$1/M \leq \lambda_{\min}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^T]) \leq \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^T]) \leq M$$

(A2) The conditional distribution $F_{Y|X}(y|x)$ is twice differentiable w.r.t. y . Denote the corresponding derivatives by $f_{Y|X}(y|x)$ and $f'_{Y|X}(y|x)$. Assume that

$$\bar{f} := \sup_{y,x} |f_{Y|X}(y|x)| < \infty, \quad \bar{f}' := \sup_{y,x} |f'_{Y|X}(y|x)| < \infty$$

uniformly in n .

(A3) $0 < f_{\min} \leq \inf_{\tau \in \mathcal{T}} \inf_x f_{Y|X}(Q(x; \tau)|x)$ uniformly in n .

A Bahadur Representation

Under Assumption (A), $m\xi_m^2 \log n = o(n)$. For any $\beta_n(\cdot)$ satisfying

$$g_n(\beta_n) := \sup_{\tau \in \mathcal{T}} \left\| \mathbb{E}[\mathbf{Z}_i \{F_{Y|X}(\mathbf{Z}_i^\top \beta_n(\tau)|X) - \tau\}] \right\| = o(\xi_m^{-1})$$

$$c_n(\beta_n) := \sup_{x, \tau \in \mathcal{T}} |Q(x; \tau) - \mathbf{Z}(x)^\top \beta_n(\tau)| = o(1)$$

we have

$$\widehat{\beta}(\tau) - \beta_n(\tau) = - \underbrace{\frac{1}{n} J_m(\tau)^{-1} \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1}\{Y_i \leq \mathbf{Z}_i^\top \mathbf{b}\} - \tau)}_{(\star)} + \text{Remainder},$$

where $J_m(\tau) := \mathbb{E}[f_{Y|X}(Q(X; \tau)) \mathbf{Z}(X) \mathbf{Z}(X)^\top]$.

Assumption (\diamond): $mc_n \log n = o(1)$, $m^3 \xi_m^2 (\log n)^3 = o(n)$, $g_n = o(n^{-1/2})$ and for any $\|\mathbf{u}\| = 1$,

$$\sup_{\tau \in \mathcal{T}} \left| \mathbf{u}^\top J_m(\tau)^{-1} \mathbb{E} \left[\mathbf{Z}_i (\mathbf{1}\{Y_i \leq Q(X_i; \tau)\} - \mathbf{1}\{Y_i \leq \mathbf{Z}_i^\top \beta_n(\tau)\}) \right] \right| = o(n^{-1/2}).$$

Under Assumption (\diamond), we can replace (\star) by

$$U_n(\tau) := n^{-1} J_m^{-1}(\tau) \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1}\{Y_i \leq Q(X_i; \tau)\} - \tau).$$

$\tau \mapsto J_m^{-1}(\tau)\mathbf{Z}_i(\mathbf{1}\{Y_i \leq Q(X_i; \tau)\} - \tau)$:

- A triangular array
- Not Lipschitz in τ

Asymptotic Equicontinuity of Quantile Process:

Under Assumption (A) and $\xi_m^2(\log n)^2 = o(n)$, we have for any $\varepsilon > 0$ and vector $\mathbf{u}_n \in \mathbb{R}^m$ with $\|\mathbf{u}_n\| = 1$,

$$\lim_{\delta \rightarrow 0} \limsup_{n \rightarrow \infty} P\left(n^{1/2} \sup_{\substack{|\tau_1 - \tau_2| \leq \delta \\ \tau_1, \tau_2 \in \mathcal{T}}} \left| \mathbf{u}_n^\top \mathbf{U}_n(\tau_1) - \mathbf{u}_n^\top \mathbf{U}_n(\tau_2) \right| > \varepsilon\right) = 0.$$

Proof: Stochastic Equicontinuity + Chaining (Kley et al., 2015; van der Vaart and Wellner, 1996)

Weak Convergence: General Series Estimator

Under Assumption (A) and (\diamond) . For a sequence \mathbf{u}_n , if

$$\begin{aligned} & H(\tau_1, \tau_2; \mathbf{u}_n) \\ & := \lim_{n \rightarrow \infty} \|\mathbf{u}_n\|^{-2} \mathbf{u}_n^\top J_m^{-1}(\tau_1) \mathbb{E}[\mathbf{Z}\mathbf{Z}^\top] J_m^{-1}(\tau_2) \mathbf{u}_n (\tau_1 \wedge \tau_2 - \tau_1 \tau_2) \end{aligned}$$

exists for any $\tau_1, \tau_2 \in \mathcal{T}$, then

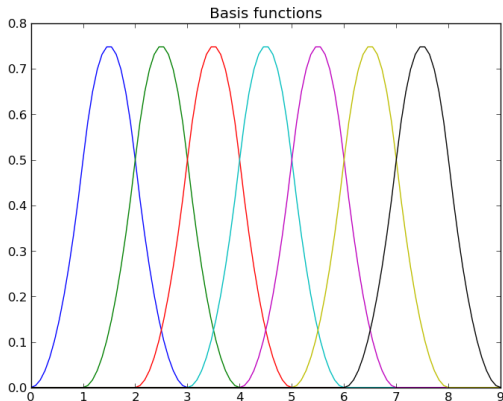
$$\frac{\sqrt{n}}{\|\mathbf{u}_n\|} \left(\mathbf{u}_n^\top \hat{\boldsymbol{\beta}}(\cdot) - \mathbf{u}_n^\top \boldsymbol{\beta}_n(\tau) \right) \rightsquigarrow \mathbb{G}(\cdot) \text{ in } \ell^\infty(\mathcal{T}), \quad (2.1)$$

where $\mathbb{G}(\cdot)$ is a centered Gaussian process with the covariance function H . In particular, there exists a version of \mathbb{G} with almost surely continuous sample paths.

Proof: Asymptotic equicontinuity, and verifying the Lindeberg condition.

Splines

B-Splines $\mathbf{B} = (b_1(x), b_2(x), \dots, b_m(x))$ are **local** basis functions



Notation: For $\mathcal{I} \in \{1, \dots, m\}$ and $\mathbf{a} \in \mathbb{R}^m$, $\mathbf{a}^{(\mathcal{I})} = (a'_j)_{j=1}^m \in \mathbb{R}^m$ where $a'_j = 0$ for $j \notin \mathcal{I}$

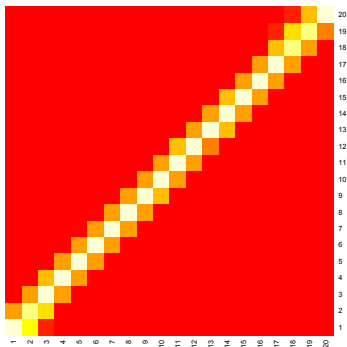
Benefit from Using Splines

$J_m(\tau)$ is a block matrix, where

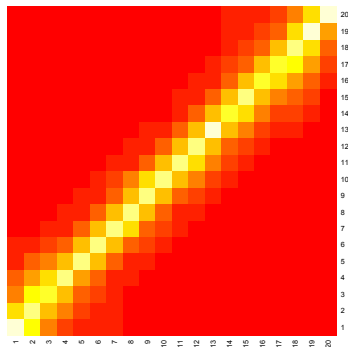
$$J_m(\tau) = \mathbb{E}[f_{Y|X}(Q(X; \tau)) \mathbf{B}(X) \mathbf{B}(X)^\top],$$

entries in $J_m^{-1}(\tau)$ decay geometrically from the main diagonal (Demko et al., 1984)

(a) $J_m(\tau)$, $\dim(X) = 1$



(b) $J_m^{-1}(\tau)$, $\dim(X) = 1$



If \mathbf{a} has at most $\|\mathbf{a}\|_0$ nonzero consecutive entries, we can find set $\mathcal{I}(\mathbf{a}) \subset \{1, \dots, m\}$ with $|\mathcal{I}(\mathbf{a})| \asymp \log n$ such that

$$\|\mathbf{a}^\top J_m^{-1}(\tau) - (\mathbf{a}^\top J_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{a}))}\| \lesssim \|\mathbf{a}\|_\infty \|\mathbf{a}\|_0 n^{-c}$$

where $c > 0$ is arbitrary

If \mathbf{u}_n is nonzero at the position in an index set \mathcal{I} , which consists of $L < \infty$ consecutive entries

$$\mathbf{u}_n^\top \mathbf{U}_n(\tau) \approx \frac{1}{n} \sum_{i=1}^n \mathbf{u}_n^\top J_m^{-1}(\tau) \mathbf{B}_i^{(\mathcal{I}(\mathbf{u}_n))} (\mathbf{1}\{Y_i \leq \mathbf{B}_i^\top \boldsymbol{\beta}_n(\tau)^{(\mathcal{I}'(\mathbf{u}_n))}\})_{-\tau}$$

where $\mathcal{I}'(\mathbf{u}_n) = \{1 \leq j \leq m : \exists i \in \mathcal{I}(\mathbf{u}_n) \text{ such that } |j - i| \leq r\}$

A reduction from dimension m to $\log n$!

Define

$$\gamma_n(\tau) := \operatorname{argmin}_{\mathbf{b} \in \mathbb{R}^m} \mathbb{E}[(\mathbf{B}^\top \mathbf{b} - Q(X; \tau))^2 f_{Y|X}(Q(X; \tau)|X)],$$

Assumption (\diamond'): $\tilde{c}_n^2 = o(n^{-1/2})$, $\xi_m^4 (\log n)^6 = o(n)$, where

$$\tilde{c}_n(\gamma_n) := \sup_{x, \tau \in \mathcal{T}} |Q(x; \tau) - \mathbf{Z}(x)^\top \boldsymbol{\beta}_n(\tau)|$$

Compare to Assumption (\diamond):

- $\xi_m = O(m^{1/2})$: $\xi_m^4 (\log n)^6 = o(n)$ is much **weaker** than $m^3 \xi_m^2 (\log n)^3 = o(n)$ in Assumption (\diamond)
- If $Q(x; \tau)$ is smooth in x for all τ , then $\tilde{c}_n = o(n^{-2/5})$ when $x \in \mathbb{R}$

Weak Convergence: Spline Series Estimator

Under Assumption (A) and (\diamond') . For a sequence \mathbf{u}_n , if

$$\begin{aligned} & \tilde{H}(\tau_1, \tau_2; \mathbf{u}_n) \\ & := \lim_{n \rightarrow \infty} \|\mathbf{u}_n\|^{-2} \mathbf{u}_n^\top J_m^{-1}(\tau_1) \mathbb{E}[\mathbf{B}\mathbf{B}^\top] J_m^{-1}(\tau_2) \mathbf{u}_n (\tau_1 \wedge \tau_2 - \tau_1 \tau_2) \end{aligned}$$

exists for any $\tau_1, \tau_2 \in \mathcal{T}$, then

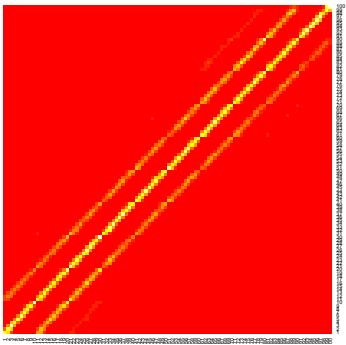
$$\frac{\sqrt{n}}{\|\mathbf{u}_n\|} \left(\mathbf{u}_n^\top \hat{\boldsymbol{\beta}}(\cdot) - \mathbf{u}_n^\top \boldsymbol{\gamma}_n(\cdot) \right) \rightsquigarrow \mathbb{G}(\cdot) \text{ in } \ell^\infty(\mathcal{T}), \quad (3.1)$$

where $\mathbb{G}(\cdot)$ is a centered Gaussian process with the covariance function \tilde{H} . In particular, there exists a version of \mathbb{G} with almost surely continuous sample paths.

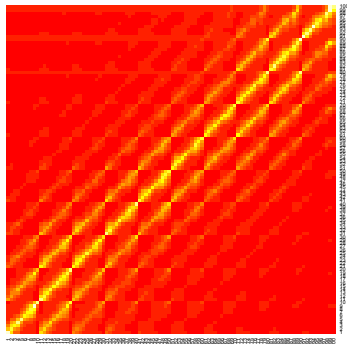
Tensor-Product Spline

Conjecture(Demko et al., 1984, Sec.5): the elements of $J_m^{-1}(\tau)$ from the tensor-product spline \mathbf{B} are also decreasing geometrically away from the main diagonal

(c) $J_m(\tau), \dim(X) = 2$



(d) $J_m^{-1}(\tau), \dim(X) = 2$



Application: Partial Linear Model

Partial Linear Model:

$$Q(X; \tau) = V^\top \boldsymbol{\alpha}(\tau) + h(W; \tau), \quad (3.2)$$

where $X = (V^\top, W^\top)^\top \in \mathbb{R}^{k+k'}$ and $k, k' \in \mathbb{N}$ are fixed.

Expanding $w \mapsto h(w; \tau)$ in terms of basis vectors $w \mapsto \tilde{\mathbf{Z}}(w)$, we can approximate (3.2):

$$Q(x; \tau) \approx \mathbf{Z}(x)^\top \boldsymbol{\beta}_n(\tau)$$

where $\mathbf{Z}(x) = (v^\top, \tilde{\mathbf{Z}}(w)^\top)^\top$ and $\boldsymbol{\beta}_n = (\boldsymbol{\alpha}(\tau)^\top, \boldsymbol{\beta}_n^\dagger(\tau)^\top)^\top$

► **Quantile Regression** gives $\hat{\boldsymbol{\alpha}}(\tau)$ and $\hat{h}(w; \tau) = \tilde{\mathbf{Z}}(w)^\top \hat{\boldsymbol{\beta}}_n^\dagger(\tau)$

Assumption (B):

(B1) Define $c_n^\dagger := \sup_{\tau, w} |\tilde{\mathbf{Z}}(w)^\top \boldsymbol{\beta}_n^\dagger(\tau) - h(w; \tau)|$ and assume that

$$\xi_m c_n^\dagger = o(1);$$

$$\sup_{\tau \in \mathcal{T}} \mathbb{E}[f_{Y|X}(Q(X; \tau)|X) \|h_{VW}(W; \tau) - A(\tau)\tilde{\mathbf{Z}}(W)\|^2] = O(\lambda_n^2)$$

$$\text{with } \xi_m \lambda_n^2 = o(1)$$

(B2) We have $\max_{j \leq k} |V_j| < C$ almost surely for some constant $C > 0$.

(B3)

$$\left(\frac{m \xi_m^{2/3} \log n}{n} \right)^{3/4} + c_n^{\dagger 2} \xi_m = o(n^{-1/2}).$$

Moreover, assume that $c_n^\dagger \lambda_n = o(n^{-1/2})$ and $m c_n^\dagger \log n = o(1)$.

Joint Process Convergence

Under Assumption (A) and (B), suppose at w_0 ,

$$\begin{aligned} & \Gamma_{22}(\tau_1, \tau_2) \\ &= \lim_{n \rightarrow \infty} \|\tilde{\mathbf{Z}}(w_0)\|^{-2} \tilde{\mathbf{Z}}(w_0)^\top M_2(\tau_1)^{-1} \mathbb{E}[\tilde{\mathbf{Z}}(W)\tilde{\mathbf{Z}}(W)^\top] M_2(\tau_2)^{-1} \tilde{\mathbf{Z}}(w_0) \end{aligned}$$

exists, then

$$\left(\begin{array}{c} \sqrt{n} \{ \hat{\boldsymbol{\alpha}}(\cdot) - \boldsymbol{\alpha}(\cdot) \} \\ \frac{\sqrt{n}}{\|\tilde{\mathbf{Z}}(w_0)\|} \{ \hat{h}(w_0; \cdot) - h(w_0; \cdot) \} \end{array} \right) \rightsquigarrow (\mathbb{G}_1(\cdot), \dots, \mathbb{G}_k(\cdot), \mathbb{G}_h(w_0; \cdot))^\top,$$

in $(\ell^\infty(\mathcal{T}))^{k+1}$ where $(\mathbb{G}_1(\cdot), \dots, \mathbb{G}_k(\cdot), \mathbb{G}_h(w_0; \cdot))$ are centered Gaussian processes with joint covariance function

$$\Gamma(\tau_1, \tau_2; \tilde{\mathbf{Z}}(w_0)) = (\tau_1 \wedge \tau_2 - \tau_1 \tau_2) \left(\begin{array}{cc} \Gamma_{11}(\tau_1, \tau_2) & \mathbf{0}_k \\ \mathbf{0}_k^\top & \Gamma_{22}(\tau_1, \tau_2) \end{array} \right) \quad \text{Detail}$$

In particular, there exists a version of $\mathbb{G}_h(w_0; \cdot)$ with almost surely continuous sample paths.

Remarks

- Most asymptotic results for partial linear model separately study the parametric and nonparametric part
- Joint asymptotics phenomenon for mean partial linear model: Cheng and Shang (2015)
- We show such joint asymptotics holds in process sense for quantile model

Thank you for your attention

arXiv 1604.02130

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Detail for PLM

$$M_2(\tau) := \mathbb{E}[\tilde{\mathbf{Z}}(W)\tilde{\mathbf{Z}}(W)^\top f_{Y|X}(Q(X;\tau)|X)]$$

$$\Gamma_{11}(\tau_1, \tau_2)$$

$$= M_{1,h}(\tau_1)^{-1} \mathbb{E}[(V - h_{VW}(W; \tau_1))(V - h_{VW}(W; \tau_2))^\top] M_{1,h}(\tau_2)^{-1}$$

where $M_{1,h}(\tau) =$

$$\mathbb{E}[(V - h_{VW}(W; \tau))(V - h_{VW}(W; \tau))^\top f_{Y|X}(Q(X; \tau)|X)]$$

► Partial Linear Model